Applications of the Poisson probability distribution

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Abstract

The Poisson distribution was introduced by Simone Denis Poisson in 1837. It has since been subject of numerous publications and practical applications. The purpose of this paper is to raise awareness of numerous application opportunities and to provide more complete case coverage of the Poisson distribution. First a formal definition and basic characteristics of a Poisson variable and its distribution are summarized. Next cases, representing time and space oriented Poisson situations, are presented. Probability assessment solutions, using functions built in spreadsheet programs, are presented. Finally, pedagogical issues, regarding challenges in teaching Statistics are discussed.

Keywords: statistics, probability, distribution, random variable, Poisson.

POISSON VARIABLE AND DISTRIBUTION

The Poisson distribution is a probability distribution of a discrete random variable that stands for the number (count) of statistically independent events, occurring within a unit of time or space (Wikipedia-Poisson, 2012), (Doane, Seward, 2010, p.232), (Sharpie, De Veaux, Velleman, 2010, p. 654), (Jaggia, S., Kelly, A., 2012, p. 157), (Donnelly, 2012, p. 215), (Anderson, Sweeney, Williams, 2012, p. 236). Given the expected value, μ , of the variable, X, the probability function is defined as:

$$f(n) = P(X = n) = \frac{e^{-u}\mu^n}{n!}, \quad F(n) = \sum_{k=0}^{k=n} f(k), \quad n = 0, 1, 2, \dots$$

A more generalized attempt to define this distribution is shown in (Levine, Stephan, Krehbiel, Berenson, 2011) where the unit of time or space is referred to as an area of opportunity (p. 177). It is an interesting generalization of this important context of the Poisson distribution.

Wolfram MathWorld (2012) arrives at this distribution by taking the limit of the Binomial (Bernoulli) distribution when the number, N, of trials becomes very large $(N \rightarrow \infty)$. A Binomial random variable represents the number of successes in a series of independent and probabilistically homogenous trials (Wikipedia-Binomial, 2012). A relation of the Poisson distribution to the Binomial distribution is also motioned in (Pelosi, Sandifer, 2003) and (De Veaux, Velleman, Bock 2006, p. 388).

Assessment of probabilities for Poisson variables is not complicated. There are many Web sites that provide calculators to this end ("Poisson Calculator ...", 2012). All major, contemporary statistical packages and spreadsheet programs are well equipped with functions for computing Poisson probabilities. For example, Google spreadsheet and Microsoft Excel provide the same syntax for the distribution functions:

$$P(X = n) = \text{Poisson}(n, \mu, false), P(X \le n) = \text{Poisson}(n, \mu, true)$$

Notice that the last parameter of the Poisson function indicates whether or not the function returns the cumulative probability. An example for computing Poisson probabilities in a Google spreadsheet is shown in Figure 1 (Appendix). A similar function is available in the Open Office Spreadsheet program, $=Poisson(x; \mu; 0|1)$. Notice that this spreadsheet program uses a semicolon to separate the function's arguments.

TIME ORIENTED POISSON VARIABLES

Feller (1966, p. 17) shows how the Poisson distribution can be derived from a series of Exponentially distributed random variables, $S_n = X_1 + X_2 + ... + X_n$. Considering a random variable, N(t), such that $N(t) = \max\{n: S_n \le t\}$ and given all variables X_k , k=1,2,...,n, have the same exponential distribution, $f(x) = \mu e^{-\mu x}$, $x \ge 0$, the N(t) variable has this distribution:

$$f(t,n) = P\{N(t) = n\} = \frac{e^{-\mu t} (\mu t)^n}{n!}, \quad F(t,n) = P\{N(t) \le n\} = \sum_{k=0}^{k=n} f(t,k), \quad n = 0,1,2,\dots$$

Consequently, the Poisson random variable can be "stretched" over longer or shorter time intervals. Since μ is the expected (average) number of events per one unit of time or space, μt

will be such a number per t units. One has to make sure that process N(t) is stationary within time interval (0,t).

Whether one observes patients arriving at an emergency room, cars driving up to a gas station, decaying radioactive atoms, bank customers coming to their bank, or shoppers being served at a cash register, the streams of such events typically follow the Poisson process. The underlying assumption is that the events are statistically independent and the rate, μ , of these events (the expected number of the events per time unit) is constant. The list of applications of the Poisson distribution is very long. To name just a few more:

- The number of soldiers of the Prussian army killed accidentally by horse kick per year (von Bortkewitsch, 1898, p. 25).
- The number of mutations on a given strand of DNA per time unit (Wikipedia-Poisson, 2012).
- The number of bankruptcies that are filed in a month (Jaggia, Kelly, 2012 p.158).
- The number of arrivals at a car wash in one hour (Anderson et al., 2012, p. 236).
- The number of network failures per day (Levine, 2010, p. 197).
- The number of file server virus infection at a data center during a 24-hour period. The number of Airbus 330 aircraft engine shutdowns per 100,000 flight hours. The number of asthma patient arrivals in a given hour at a walk-in clinic (Doane, Seward, 2010, p. 232).
- The number of hungry persons entering McDonald's restaurant. The number of workrelated accidents over a given production time, The number of birth, deaths, marriages, divorces, suicides, and homicides over a given period of time (Weiers, 2008, p. 187).
- The number of customers who call to complain about a service problem per month (Donnelly, Jr., 2012, p. 215).
- The number of visitors to a Web site per minute (Sharpie, De Veaux, Velleman, 2010, p. 654).
- The number of calls to consumer hot line in a 5-minute period (Pelosi, Sandifer, 2003, p. D1).
- The number of telephone calls per minute in a small business. The number of arrivals at a turnpike tollbooth par minute between 3 A.M. and 4 A.M. in January on the Kansas Turnpike (Black, 2012, p. 161).

Consider a simple emergency room example where 2 patients arrive, on average, every 10 minutes (this is equivalent to 0.2 patients per one minute). Consecutive arrivals are statistically independent. This means that each given arrival has no impact on the probability of next arrivals. Letting N(t) represent the number of such arrival in t minutes, one can evaluate the following probabilities, using the Poisson distribution (as implemented in Excel or Google spreadsheet):

 $P\{N(60) = 10\} = Poisson(10,60*0.2,False) = 10.48\%$ $P\{N(60) \le 10\} = Poisson(10,60*0.2,True) = 34.72\%$ $P\{N(60) < 10\} = Poisson(9,60*0.2,True) = 24.24\%$ $P\{N(60) > 10\} = 1-Poisson(10,60*0.2,True) = 65.28\%$ $P\{N(60) \ge 10\} = 1-Poisson(9,60*0.2,True) = 75.76\%$ Notice that in real world situations, due to the aspect of "seasonality", the arrival rate may remain constant only within limited time intervals ("seasons"). In such situations, variable *t* must not go beyond the upper limits of such seasonal intervals. For example, a morning rate may stay constant between 5:00 and 11:00 a.m., a noon rate— between 1:00 a.m. and 2:00 p.m., etc. Figure 2 (Appendix) shows implementation of the above formulas in a Google spreadsheet.

Time based Poisson variables are more popular. One can find numerous examples and cases, involving such random variables in most to the contemporary statistical textbooks or papers.

SPACE ORIENTED POISSON VARIABLES

Space oriented Poisson variables are less common or popular. Nonetheless, one can find some interesting examples of such variables in most recent Statistics books. For example:

- The number of typographical errors found in a manuscript or the total number of home runs hit in Major League Baseball games (Donnelly, Jr., 2012, p. 215).
- The number of surface defects on a new refrigerator or the number of fleas on the body of a dog (Levine et al., 2011, p.).
- The number of blemishes per sheet of white bond paper (Doane, Seward, 2010, p. 232).
- The number of repairs needed in 10 miles of highway or the number of leaks per 100 miles of pipeline (Anderson et al., 2012, p. 237).
- The number of defects in a 50-yard roll of fabric or the number of bacteria in a specified culture (Jaggia, Kelly, 2012, p.158).
- The number of defects that occur in a computer monitor of a certain size (Sharpie et al., 2010, p. 654).
- The number of a certain type of insect that can be found in a 1-square-foot area of farmland (Pelosi, Sandifer, 2003, p. D1).
- The number of hazardous waste sites per county in the United States or the number of sewing flaws per pair of jeans during production (Black, 2012, p. 161).
- The number of eagles nesting in a region or the number of hits by V-1 buzz bombs in WWII London (Triola, 2007, p. 252-253).

Among the above cases, of a particular interest is the last one. During the World War II, Germans fired thousands of so called V-1 buzz bombs over the English channel toward London. About 800 bombs managed to hit some targets in London. The rest either did not make it over the water or were destroyed by RAF. A particular case, presented by Triola (2007, p. 253), focuses on South London. This region was subdivided into 576 regions, each of 0.25 km². The entire region was hit by 535 V-1 bombs.

By analogy to the number of patients arriving at an emergency room per one minute, the random variable for the V-1 bomb case, represents the number of bomb hits per region. Since the each area is of the same size, the average number of hits par one area can be expressed as: $\mu = 537/575 = 0.9288$. Figure 3 shows a Google spreadsheet based model for this case. The probabilities and expected number of areas hit the x V1 bombs are calculated in the following way:

P(X=n) = Poisson(n, 0.9288, False)

 $E(\text{AreaCount} \mid n) = 576*\text{Poisson}(n, 0.9288, \text{False})$

Detail spreadsheet solution is shown in (Google-V1 Bombs, 2012).

Feller (1961, p. 145) shows remarkably close fit of the actual numbers of bomb hits as compared to the expected numbers provided by the Poisson distribution. Figure 3 (Appendix) shows the numbers along with a goodness-of-fit test outcomes.

APPROXIMATION OF THE BINOMIAL DISTRIBUTION

As shown above, the Poisson distribution is a special case of the Binomial distribution. In some situations for former one can be used to approximate the latter one. It is particularly feasible if, for of a Binomial random variable, the number of trials, n, is extremely large and the probability of success, p, is very small. According to Triola (2007, p. 254) the Poisson distribution provides a good approximation of the Binomial distribution, if $n \ge 100$, and $np \le 10$. In such situations, events attributed to successes are called rare events. The Poisson distribution has been particularly useful in handling such events.

A leukemia case of Woburn, Massachusetts, falls into the category of rare-event situations De Veaux et al. (2006 p. 387-388) shows the following case:

In early 1990s, a leukemia cluster was identified in the Massachusetts town of Woburn. Many more cases of leukemia, a malignant cancer that originates in a cell in the marrow of bone, appeared in this small town than would be predicted. Was it evidence of a problem in the town, or was it a chance? That question led to a famous trial in which the families of eight leukemia victims sued and became grist for a book and movie *A Civil Action*. Following an 80-day trial, the judge called for a retrial after dismissing the jury's contradictory and confusing findings. Shortly thereafter, the chemical companies and the families settled.

The issue of evidence versus chance is at a core of statistical studies. Using common data for the same period, one can estimate the probably of leukemia cases nationwide and then compare them with local results. The total US population of 280,000,000 and an annual average of leukemia cases of 30,8000 provide the Binomial probability, p, of success of about 30,8000/280,000,000 =0.00011. Detail calculations are shown in (De Veaux et al., 2006 p. 388-389) and in a spreadsheet at Google-Woburn (2012). Figures 4 and 5 (Appendix) show the output. With the Woburn population of 35,000, the probability of at least 8 leukemia cases, as computed in at Google-Woburn (2012), using the Binomial and Poisson functions, is:

 $P(X \ge 8) = 1$ - BINOMDIST(7,35000, 0.00011, True) = #NUM

 $P(X \ge 8) = 1 - POISSON(7,35000*0.00011,True) = 4.25\%$

So the Binomial function in a Google spreadsheet fails to compute this probability, reporting that "argument 35000" is too large, as shown in Figure 4 (Appendix). On the other hand, the Poisson function provides a satisfactory approximation. Interestingly, in an Excel spreadsheet both the Binomial and Poisson function provide the same [correct] outcome, as shown in Figure 5 (Appendix).

The assessed probability of 4.25% is small and can be attributed to sources other than chance. Some may consider a too-close-to-call value. It is below but quite close to the common

significance level of 5%. As De Veaux et al. (2006 p. 388) concludes: "That's small but not terribly unusual."

CONCLUSIONS

The Poisson distribution has a strong theoretical background and very wide spectrum of practical applications. Bringing original and/or unusual cases, featuring Poisson processes, may provide opportunities for increasing students' attentiveness and interests in Statistics. One important lesson the author of this paper has learned is that presentation of statistical cases, including Poisson examples, should be accompanied and enriched by significant business, social or historical background description and discussion. For example, a presentation of the WWII V1 London bombing case may be accompanied by a short youtube.com movie, showing how V1 Buzz bombs were being launched to hit London targets (V- 1 Flying Bomb, 2011). Such an exposure to background information makes "technical" cases more interesting and instructive.

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Appendix

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| 2 | | Exploring Poisson Distribution | | | | | | | | | |
| 3 | | | | | | | | | | | |
| 4 | | X | | | a Poisson random variable | | | | | | |
| 5 | | | | | | | | | | | |
| 6 | | μ | 10 | | the expected n | umber of events per | one unit of time or sp | ace | | | |
| 7 | | | | | | | | | | | |
| 8 | | n | 15 | | the number of t | the number of the events to occur within one unit of time or space | | | | | |
| 9 | | P(X = 15) | 3.47% | | the probability that exactly 15 events will occur | | | | | | |
| 10 | | P(X ≤ 15) | 95.13% | | the probability that up to (no more than) 15 events will occur | | | | | | |
| 11 | | | | | | | | | | | |
| 12 | | n1 | 8 | | at least 8 | overte will occur | | | | | |
| 13 | | n2 | 12 | | no more 12 | events will occur | | | | | |
| 14 | | P(8≤X≤12) | 57.13% | | the probability that between 8 and 12 events will occur | | | | | | |

Figure 1 Doing the Poisson distribution in a Google spreadsheet (Google-Poisson, 2012).

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| | Exploring Poisson Distribution - Case: Emergency Room | | | | | | | | | | |
| | | | | | | | | | | | |
| | N(t) | | | the number of patients arriving in t minutes | | | | | | | |
| | | | | | | | | | | | |
| | μ | 0.2 | | the expected number of events per one minute | | | | | | | |
| | t | 60 | | the length of the time interval | | | | | | | |
| | n | 10 | | the number of patients to arrive in 60 minutes | | | | | | | |
| | P{N(60) = 10} | 10.48% | | the probability that exactly 10 patients will arrive in 60 minutes | | | | | | | |
| | P{N(60) ≤ 10} | 34.72% | | the probability that | up to (no more than) [,] | 10 patients will arrive | in 60 minutes | | | | |
| | P{N(60) < 10} | 24.24% | | the probability that | fewer than 10 patients | will arrive in 60 minu | utes | | | | |
| | P{N(60) > 10} | 65.28% | | the probability that more than 10 patients will arrive in 60 minutes | | | | | | | |
| | P{N(60) ≥ 10} | 75.76% | | the probability that at least 10 patients will arrive in 60 minutes | | | | | | | |
| | | | | | | | | | | | |
| | n1 | 8 | | at least 8 | antianta will arrive in 60 minutes | | | | | | |
| | n2 | 12 | | no more than 12 | patients will arrive in ou minutes | | | | | | |
| | P(8≤X≤12) | 48.65% | | the probability that between 8 and 12 patients will arrive in 60 minutes | | | | | | | |
| | | =1-POISSON(C8-1,C6*C7 A B A B N(t) µ t P{N(60) = 10} P{N(60) ≤ 10} P{N(60) < 10} P{N(60) > 10} P{N(60) ≥ 10} | Image: Point of the second secon | Image: Section of the section of t | S%123 -12pt \div BAbs:AAA=1-POISSON(C8-1,C6°C7,true)ABCDECExploring Poisson Distribution -N(t)the number of patieN(t)the number of patieP{N(t)the expected numbt60the number of patieP{N(60) = 10}10.48%P{N(60) < 10}24.24%P{N(60) > 10}65.28%the probability thatP{N(60) > 10}75.76%the probability thatP{N 8 ≥ 12}48.65%the probability that | Image: Solution of the second seco | Image: Solution of the state of the stat | | | | |

Figure 2 The Poisson distribution. An emergency room situation (Google-Emergency Room, 2012).

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| 3 | | | | | | _ | | | | | | |
| 4 | | Х | | | a Poisson random v | son random variable: the number of V1 bombs hitting a region. | | | | | | |
| 5 | | bomCnt | 537 | | the total number of | V1 bombs that hit S | outh Loandon | - | | | | |
| 6 | | areaCnt | 576 | | the number of areas | (subdivisions) of So | outh London | | | | | |
| 7 | | μ | 0.9323 | | the expected number | er of V1 bomb hits p | er area | | | | | |
| 8 | | | | | | | | | | | | |
| 9 | | n | 2 | | the number of the be | ombs to hit a randon | nly selected region | | | | | |
| 10 | | P(X = 2) | 17.11% | | the probability that e | exactly 2 events will | occur | | | | | |
| 11 | E | (areaCnt n=2) | 98.54 | | the expected number | er of areas hit 2 time | s | | | | | |
| 12 | | | | | and expected human | | | | | | | |
| 13 | | | | | | | | | | | | |
| 14 | | | | Observe | d vs. Expected Area | Count | | | | | | |
| 15 | | | | x | f(x) | Expected Number of Areas Hit by x Bombs | Observed Number of Areas Hit by x Bombs | Squared Difference | х2 | | | |
| 16 | | | | 0 | 0.3937 | 226.7 | 229 | 5.2900 | 0.0233 | | | |
| 17 | | | | 1 | 0.3670 | 211.4 | 211 | 0.1600 | 0.0008 | | | |
| 18 | | | | 2 | 0.1711 | 98.5 | 93 | 30.2500 | 0.3071 | | | |
| 19 | | | | 3 | 0.0532 | 30.6 | 35 | 19.3600 | 0.6327 | | | |
| 20 | | | | 4 | 0.0124 | 7.1 | 7 | 0.0100 | 0.0014 | | | |
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| 20 | | | | 1.1903 | | peneer m. | | | | | | |

Figure 3 The Poisson distribution. A V1 buzz WWII London bombing situation (Google-V1 Bombs, 2012).

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| 2 | | | | Bi | nomial Distributi | on | | | | | |
| 3 | | | | | | | | | | | |
| 4 | | N | N 280,000,000 US population | | | | | | | | |
| 5 | | Nl | 30,800 | | the total number of leukemia cases per year | | | | | | |
| 6 | | р | 0.00011 | | the probability that a randomly selected person has leukemia | | | | | | |
| 7 | | n | 35,000 | | Woburn, MA population | | | | | | |
| 8 | | x | 8 | | the number of leuker | mia case | s | | | | |
| 9 | | P(X = 8) | #NUM! | erro | r: Argument too large | : 35000 | kemia cas | es in Woburn, MA | | | |
| 10 | | P(X ≥ 8) | #NUM! | | | | | | | | |
| 11 | | | | | | | | | | | |
| 12 | | | | Po | oisson Distributio | on | | | | | |
| 13 | | | | | | | | | | | |
| 14 | | μ | 3.85 | | the expected number of leukemia cases in Woburn, MA | | | | | | |
| 15 | | P(X = 8) | 2.55% | | the probability of exactly 8 leukemia cases in Woburn, MA | | | | | | |
| 16 | | P(X ≥ 8) | 4.27% | | the probability of at least 8 leukemia cases in Woburn, MA | | | | | | |

Figure 4 The Binomial and Poisson distributions. A Woburn, MA leukemia situation (Google-Woburn, 2012).

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| 3 | | | | | | | | | | | |
| 4 | | N | 280,000,000 | | US population | I | | | | | |
| 5 | | Nl | 30,800 | | the total numb | per of le | ukemia cases | s per ye | ear | | |
| 6 | | р | 0.00011 | | the probability that a randomly selected person has leukemi | | | | | | |
| 7 | | n | 35,000 | | Woburn, MA population | | | | | | |
| 8 | | x | 8 | | the number of | leuken | nia cases | | | | |
| 9 | | P(X = 8) | 2.55% | | the probability | of exa | ctly 8 leukemi | ia case | s in Wo | burn, MA | |
| 10 | | P(X ≥ 8) | 4.27% | | the probability | of at le | east 8 leukem | ia case | es in Wo | oburn, MA | |
| 11 | | | | | | | | | | | |
| 12 | | | | F | oisson Dis | tribut | ion | | | | |
| 13 | | | | | | | | | | | |
| 14 | | μ | 3.85 | | the expected | number | r of leukemia o | cases i | n Wobu | rn, MA | |
| 15 | | P(X = 8) | 2.55% | | the probability | of exa | ctly 8 leukemi | ia case | s in Wo | burn, MA | |
| 16 | | P(X ≥ 8) | 4.27% | | the probability | of at le | east 8 leukem | ia case | s in Wo | oburn, MA | |

Figure 5 The Binomial and Poisson distributions. A Woburn, MA leukemia situation (an Excel workbook).