Do declining exchange rates help the U.S. economy?

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ABSTRACT

Falling exchange rates cause a decline in the value of the U.S. dollar, increasing import prices and reducing export prices. These relative price changes create income and substitution effects. This study uses an econometric model of the U.S. economy, developed using 1960-2000 data, to assess the relationship of recent changes in the exchange rate to changes in demand for imports and exports. The study finds the positive effects of export growth and substitution toward domestic goods, more than offset the negative effects on real income of higher import prices resulting from the 2000-2008 decline in U.S. exchange rate. The paper estimates the declining exchange rate (16.4%) stimulated exports and domestic consumer demand sufficiently to add 3.7% to total GDP growth over the period, and significantly decreased the trade deficit ($187.2 billion) from what it otherwise would have been. This decline in the trade deficit and the increase in domestic savings from higher incomes increased U.S. owned assets by an estimated $114 billion compared to what they would otherwise have been. JEL E00, F40, F43.

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1. Introduction

A decline in the U.S. exchange rate (XR) decreases the amount of foreign currency a dollar can buy, which can increase import prices. Such a decline occurred during the 2000 – 2008 period, making foreign goods Americans purchase more expensive, thereby reducing American real incomes. This “income effect” may reduce U. S. demand for both domestic and imported goods. It may also cause a “substitution effect” by making imports more expensive: demand may shift toward cheaper American goods. Also, the cheaper U.S. dollar may make U.S. goods cheaper, increasing American exports.

While all of these effects may occur to some extent or another, the real question is whether exchange rate changes bring about major differences in the demand for domestic goods and imports? Or perhaps even major changes in America’s exchange rate only bring about marginal, relatively insignificant, total effects. This may be because some individual effects offset others, or are only of a small magnitude relative to the GDP.

These are empirical questions. This study attempts to answer them using estimates of exchange rate changes impact on U.S. demand for both domestic and imported goods during that period 1960-2000, using an econometric model of the U.S. economy patterned, in abbreviated fashion, after the larger demand – driven econometric models of Ray Fair (2004). Fair’s model develops estimates of the GDP based on estimates of consumer and investment demand, and import demand.

These model’s findings’ as to the marginal effects of different variables are then used to simulate the effects of exchange rate changes during the 2000-2008 period. Econometric estimates of the marginal impact of the exchange rate on U.S. demand for imports and domestic goods, controlling for important variables that affect consumer and investment demand, developed by Heim (2008c) are used. These estimates, in turn, were developed from econometric analysis by Heim (2008 a, b) to determine which other variables (and which lagged values of those variables) were significant determinants of demand. The model is used as a basis for simulating the initial impact of a change in exchange rates on the GDP. This done using the exchange rate variable’s regression coefficients in the consumption, investment and export equations of the model. The initial impact is then augmented by estimates of multiplier/accelerator effects derived from the same econometric model. Three separate methods for calculating these effects are presented. Each reaches the same conclusions about the effects of the 2000 - 2008 exchange rate changes on the U.S. economy, when the U.S. Federal Reserve’s Nominal Broad trade weighted exchange rate index declined 16.4%, from 119.45 to 99.83.

Finally, the paper also estimates the extent to which exchange rate - reduces the trade deficit and reduces transfers of U.S. capital to other nations or their citizens, to pay for trade deficits.

2. Methodology

This study examines how declining exchange rates 2000 – 2008 have affected U.S. demand for domestic and imported goods. It uses a seven behavioral equation model of the U.S. economy (three consumer demand equations, three investment demand equations and an export demand equation) to estimate the GDP and its components. The three consumption equations are for total, domestically produced and imported consumer goods, the three investment equations
are for total, domestically produced and imported investment goods. The seventh equation estimated export demand.

The econometric approach, is patterned, after the more detailed (30 behavioral equations) demand – driven econometric models of Ray Fair (2004). Fair, for example, has four separate behavioral equations for household demand; whereas the model used in this study has two consumption equations: one for domestically produced consumer goods and one for imports. Fair’s model, as does this model, estimates the GDP additively, from behavioral equation estimates of consumer, investment, export and import demand. In both models, government spending on goods & services is treated as exogenous. Finally, like Fair’s model, the model here is Keynesian i.e., demand driven in orientation.

There are some differences between the models aside from size. All imports in Fair’s model are estimated as one variable and imports are modeled as simple functions of GDP growth alone. In the model used here, consumer and investment imports are modeled separately and as functions of the same large number of specific determinants of consumer and investment demand found to be important determinants of demand for domestically produced consumer and investment goods, such as wealth, profits, interest rates, depreciation, credit crowd out, etc. In Fair’s model exports are exogenous, but in the model used here exports are endogenous. They are determined by the exchange rate and a proxy for our trading partners’ economic growth rate. Another difference is that Fair commonly uses lagged values of an equation’s dependent variable on the right hand side of an equation to explain the movement in the dependent variable; the model used here does not. Its main objective is to explain the past influence of specific variables, especially the exchange rate, on consumption, investment and the GDP. Lagged values of dependent variables may improve the accuracy of predictions, which is a core objective of Fair’s model, but tend to hide from us the underlying variables that drive them (as well as the current dependent variable). Hence, they provide an inadequate explanation of underlying structural relationships. That said, in quarterly data models, lags may be needed simply to capture lagged adjustment effects. The annual data used in this model reduces that need appreciably.

Also because of Fair’s findings, equations in this model do not include variables to account for rational expectations – driven behavior, since Fair, like others before him, found little support for these issues in extensive tests in his own models, (Fernandez-Villaverde, 2008). Fair also found his own model (a Cowles – Commission type model, like the model used here) performed better than a VAR model against which he tested it, and generally better in tests against an autoregressive components (AC) model. (Fernandez-Villaverde, 2008).

A significant difference between this study and Fair’s models is the way in which autocorrelation is treated. Generally, here it is dealt with by first differencing data. In Fair, it is dealt with by leaving the data in levels and using standard autocorrelation control AR(i) variables. Generally, though not always, the first differencing used here was successful in bringing Durban Watson statistics up to acceptable levels. This approach also provided two critically important additional benefits:

1. First differencing significantly reduced multicollinearity between the variables thought to be determinants of consumption or investment. This provided for much more stable regression coefficients on variables in the model when changes to the model were made, and therefore, more stable estimates of marginal impact.

2. First differencing eliminates the irrational tendency for the regression coefficients on a particular variable (e.g., the exchange rate) in imports and domestic goods demand

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equations not to add up to same variable’s coefficient in total demand equation when using standard AR(i) controls. The two parts together (demand for imports and demand for domestically produced goods) definitionally equal total consumption or investment. Adding the estimated effects of a particular variable on import demand and domestic demand should tell us precisely how total demand is affected. Statistical results should yield the same result, assuming regression does not provide illogical results. Statistical results do equal the arithmetic sum of these two parts, unless AR(i) controls are used with any of the equations.

For example, equations 9 -11 below provide statistical estimates of the impact of the exchange rate (and other variables) on demand for domestically produced, imported and total investment goods. Equation eleven’s relatively low Durbin Watson statistic indicates possible autocorrelation. Using a standard AR(1) control raises the Durbin – Watson statistic to more acceptable levels, but at the price of changing all of other regression coefficients in that equation so that the coefficients on a variable in the domestic and imports demand equations no longer add to the total effect, as they did – exactly - before the autocorrelation control was added!. The situation is not improved by adding autocorrelation controls to the other equations.

To avoid this problem, first differencing is used here as the method of autocorrelation control. Where first differencing is not successful in raising the Durbin Watson statistic to non-autocorrelation levels, the evidence suggests the coefficient is not adversely affected, hence we ignore it. For example, subtracting the (non-) autocorrelation plagued imports coefficient (eq. 10), from the (non-) autocorrelation plagued total investment coefficient for the same variable (eq. 9) yields exactly the value of a variable we find in the autocorrelation plagued domestic demand equation (eq. 11).

Arithmetic methods for estimating the effect of exchange rate changes on the GDP, consumption, investment, the trade deficit, and U.S. ownership of assets are developed and presented in sections 9-12 of this study. They depend heavily on the marginal effects of changes in exchange rates given by regression coefficients in our econometric models. The reliability of such estimates is greatly enhanced if the models have made serious efforts to control for all other major variables that effect consumption and investment. This provides better assurance that the estimated effects of the exchange rate are good estimates, and not estimates which mislead because they are also proxying for the effects of other consumption or investment determinants not controlled for in the testing process. Extensive efforts were made in Heim (2008a, b &c) to determine what theoretically – postulated variables belonged in these equations, and which lagged value of the variables was the most appropriate to use. The consumption and investment models used here utilize the findings from those studies coupled with the Federal Reserve’s (Nominal) Broad exchange rate. They are described further below.

All data used in those studies was taken from the Council of Economic Advisors’ statistical appendix to the Economic Report of the President, 2002. Data Tables B2, B3, B7, B26, B54, B60, B73, B82, B90, B95, B104, B106 and B110. However, additional multilateral trade weighted value of the dollar, i.e., the foreign exchange rate data, is taken from Table B110 of the Economic Report of the President, 2001 and Table B108 of the 1997 Economic Report of the President, 1997. Exchange rate values 1960 - 1970 were assumed constant at 1970 levels, per the Bretton Woods protocols. All data are expressed in real 1996 dollars, or converted to same using the GDP deflator in Table B3.
Each regression below shows the estimated marginal effect (regression coefficient) for the explanatory variables, the t statistic associated with it, the percent of variance explained and the Durbin Watson autocorrelation statistic. Depending on the particular regression test and the number of lags used, our sample size was 36-38 observations from the 1960-2000 period. With this number of observations, throughout the remainder of the paper, marginal effects with a t-statistic of 1.8 are significant at the 8% level, 2.0 are significant at the 5% level and t-statistics of 2.7 are significant at the 1% level.

Because of the simultaneity between the total consumption variable (C) in the GDP accounts, or its component part, domestic consumer goods (C_D), and income (Y) inherent in these equations, two stage least squares estimates of disposable income $\Delta (Y-T_G)_0$ were used. The remaining right hand side variables were used as first stage regressors. Newey-West heteroskedasticity corrections were also made, generally improving t-statistics. Two Stage least Squares was also used with the investment equations because of simultaneity between investment and the economy’s growth rate (the accelerator variable).

There is some difficulty separating consumer imports out of total imports in the *Economic Report of the President*. This is because, as the Bureau of Economic Analysis (BEA) has confirmed, it does not categorize import and export data into same “C” and “I” and “G” categories used elsewhere in the national GDP accounts. Absent official determinations by BEA, economists must make their own evaluations of how to divide the data. It is not clear from Table 104 in the *Economic Report of the President*, for example, how much of the value of motor vehicle imports or petroleum imports are for business (inventory investment) vs. consumer use. Data on imported services (Table B-106) does not distinguish between imports of services by businesses and consumers, though one might suspect the former dominate. Nor do the services data extend back beyond 1974, so no deduction from total imports for business services imports could be made in calculating consumer imports.

Following Heim (2007), we then take as our definition of consumer goods and services imports all imports except for imports of capital goods and industrial supplies and materials. The theory behind this choice was that the best definition of “consumer” imports was the one whose variation was best explained (highest $R^2$) by the variables theoretically thought to drive demand for consumer imports. Other definitions of consumer imports, did not explain consumer behavior as well and were rejected.

Hence, for consumer imports, the definition used is

$$(M_{m-ksm}) = Total\ Imports\ (M) - (Capital\ Goods\ Imports + Imported\ Industrial\ Supplies\ and\ Materials(M_{ksm}))$$

These definitions appear to be reasonable, if not exact, given the data available. Separate regressions were then run on total consumer demand, and separately for imported consumer goods alone. Results for the imports equation were subtracted from the results for the total consumption (C) equation, to estimate demand for domestically produced consumer goods. As noted earlier when discussing autocorrelation, the coefficients obtained in this manner (arithmetically) for each variable are exactly the same as those obtained statistically by regressing these same determinants on domestically produced consumer goods (C-$M_{m-ksm}$).

Investment imports were defined using the same process as imports of capital goods plus imports of industrial supplies and materials ($M_{ksm}$), i.e., total imports minus consumer imports.
Preliminary testing suggested that exchange rates have some lagged effects that go back as far as three years ago, so the average exchange rate for those years ($XR_{Av0123}$) was used. Individual variables for each year’s exchange rate were not used. High levels of multicollinearity between the individual years’ exchange rates made coefficient values for any one year change dramatically when another year’s exchange rate variable was added or deleted. However, the coefficients on the average exchange rate variables tended to precisely or approximately add up to the sum of the coefficients when separate exchange rate variables were used for each year. In addition, adding an additional year’s lag to the average increased explained variance, up through the three year lag. This suggests that the full effects of exchange rate changes take that long to achieve. For example, peoples’ demand may be conditioned on what they recall price has been in the recent past as well as what it is today. It may also be that there are long lead times required for delivery of some items, e.g., machinery. If so, this year’s actual purchases may have been the result of a prior year’s decision to purchase, based on a prior year’s price determined in part by that year’s exchange rate.

In sections 9 below the results to the economic system attributable to a shock to exchange rates are used to estimate the resulting effects on the GDP, consumption and investment demand (both domestic and imports). Three separate methods for calculating these results are used. They yielded identical results:

1. The first of the three methods simply sums the three regression coefficient estimates of the effect of an exchange rate change on (1) domestically produced consumer goods and (2) investment goods, and (3) exports and uses the estimated multiplier/accelerator effect developed in Section 8 from the same econometric model to calculate the total effect on income and consumption.

2. A second method, commonly used in large scale econometric models such as Fair’s to calculate the effects of a shock was also used as a check on the first. This method involves estimating the consumption, investment and import equations from their determinants and combining them with exogenous information on export demand and government demand to (additively) calculate the GDP.

3. A Keynesian “IS” curve model calculated directly from our econometric results. This IS model is used to predict changes, modified by the model’s estimated multiplier effect, in the GDP likely to occur when the exchange rate shock occurs.

Finally we note that Heim (2008c) uses these same econometric models for investment and consumption demand as used in this study. However, it utilizes a less satisfactory, more ambiguous method for separating the income effects of exchange rate changes from substitution effects, and overstates both income and substitution effects as a result. Since then, a more rigorous method for isolating the separate income and substitution effects has been developed (see sections 8 and 9 below), which markedly changes the estimates of the GDP, trade deficit, etc. obtained in Heim 2008c). There are other differences as well: this study expands estimates of effects to cover to cover the 2008 period, and uses substantially changed regression estimates in the export model.

3. The Consumer Demand Model:

To study the effect of exchange rate changes on consumption we need a theory of consumer demand, so that in testing, we can control for changes in consumption causes by things
other than the exchange rate itself. This paper postulates a modified Keynesian theory of demand for consumer goods. It assumes that in general, the determinants of the demand for both domestic and imported consumer goods are the same as those mentioned in Keynes (1936), with the addition of two other variables. First, a “crowd out” variable is added, similar to the one used in investment studies to control for periods of limited credit availability which may occur in response to government deficits. Second, we also add an exchange rate variable.

Keynes argues in chapter 8 of the *General Theory of Employment, Interest and Money* (1936) that income, wealth, fiscal policy (taxes) and possibly the rate of interest might influence consumption. However, he felt

\[ \text{income}\ldots \text{is, as a rule, the principal variable upon which the consumption-constituent of the aggregate demand function will depend…windfall changes in capital-values will be capable of changing the propensity to consume, and substantial changes in the rate of interest and in fiscal policy may make some difference (pp.95-96)…} \]

where “fiscal policy” is a reference to tax levels. In chapter 9 he also notes other factors that might affect the level of consumption spending: precautionary saving (for unknown, but potential, future needs), saving for known future needs (like retirement), and saving to finance improvements in future standards of living.

Heim (2008b) found that regression results on a modified Keynesian function of the following type explained about 90% of the variance in consumer spending in the 1960 - 2000 period:

\[
C = \beta_1 + \beta_2 (Y-T_G) + \beta_3 (T_G - G) - \beta_4 (PR) + \beta_5 (DJ) - 2 + \beta_6 (XR)_{AV0123}
\]  

where

\[
(Y-T_G) = \text{Total income minus taxes, defined as the GDP minus the portion of total government receipts used to finance government purchases of goods and services, i.e., total government receipts minus the portion used to finance government spending on transfer payments not included in the GDP definition of government spending.}
\]

\[
(T_G - G) = \text{The government deficit, interpreted as a restrictor of consumer as well as investment credit. Usually we will disaggregate this into two separate variables in regressions: } \beta_{3A} T_G(0) \text{ and } \beta_{3B} G. \text{ because it has been found the effects of each on consumer spending differs, with the tax variable the more important. (Heim 2008a)}
\]

\[
PR = \text{An interest rate measure, the Prime rate, for the current period. This rate is a base rate for much consumer credit. It is deflated to get the “real” rate using the average of the past two year’s CPI inflation rate.}
\]

\[
DJ_{-2} = \text{A stock market wealth measure, the Dow Jones Composite Average, lagged two years}
\]
The trade-weighted exchange rate (XR), averaged over four years. In our regressions, an average of the XR value for the current and past three years is used, denoted XR_{AV0123}. This is done to capture what preliminary studies showed was slow, multiyear process of adjustment to exchange rate changes.

First difference versions, shown below, of this consumption function (1) were used to reduce the distorting effects of multicollinearity and non-stationarity inherent in most time series econometric models:

\[ \Delta C_0 = \beta_2 \Delta (Y-T_G)_0 + \beta_3 \Delta (T_G - G)_0 - \beta_4 \Delta (PR)_0 + \beta_5 \Delta (DJ)_{-2} + \beta_6 \Delta (XR)_{AV0123} \]

or

\[ \Delta C_0 = \beta_2 \Delta (Y-T_G)_0 + \beta_3A \Delta (T)_G(0) - \beta_3B \Delta (G)_0 - \beta_4 \Delta (PR)_0 + \beta_5 \Delta (DJ)_{-2} + \beta_6 \Delta (XR)_{AV0123} \]

These last two equations are the same except that in (3) we have divided the crowd out variable into two variables. We will test these hypotheses, particularly (3), further below, and use the results to calculate the effects of exchange rate change on consumer demand.

Using the nominal broad exchange rate, the government deficit variables, and the Keynesian variables, our regression findings for consumer demand model are as follows:

**Total Consumer Demand**

\[ \frac{\Delta C_0}{\Delta (Y-T_G)_0} = 0.63 \Delta (Y-T_G)_0 + 0.47 \Delta (T_G)_0 + 0.06 \Delta G_0 - 6.22 \Delta PR_0 + 0.60 \Delta DJ_{-2} + 2.69 \Delta XR_{AV0123} \]

\[ R^2 = 91\% \]

\[ (t) \]

\[ (18.2) \quad (4.3) \quad (0.5) \quad (-3.4) \quad (4.0) \quad (1.5) \]

\[ \text{D.W.} = 1.7 \]

**Demand for Imported Consumer Goods**

\[ \frac{\Delta (C-M_{m-ksm})_0}{\Delta (Y-T_G)_0} = 0.06 \Delta (Y-T_G)_0 + 0.27 \Delta (T_G)_0 - 0.18 \Delta G_0 - 3.94 \Delta PR_0 + 0.26 \Delta DJ_{-2} + 4.33 \Delta XR_{AV0123} \]

\[ R^2 = 83\% \]

\[ (t) \]

\[ (2.8) \quad (4.6) \quad (-1.9) \quad (-2.6) \quad (1.8) \quad (2.9) \]

\[ \text{D.W.} = 1.5 \]

**Demand for Domestically Produced Consumer Goods**

\[ \frac{\Delta (C-M_{m-ksm})_0}{\Delta (Y-T_G)_0} = 0.57 \Delta (Y-T_G)_0 + 0.20 \Delta (T_G)_0 + 0.24 \Delta G_0 - 2.28 \Delta PR_0 + 0.34 \Delta DJ_{-2} - 1.64 \Delta XR_{AV0123} \]

\[ R^2 = 74\% \]

\[ (t) \]

\[ (16.0) \quad (1.6) \quad (1.3) \quad (-0.7) \quad (2.8) \quad (-1.0) \]

\[ \text{D.W.} = 1.8 \]

Though not presented here, the same models without the exchange rate variable had \( R^2 \) of 91, 77 and 74% respectively. The exchange rate appears to have a major influence on import demand, adding 6% points to explanatory power, but seems to have a minimal effect on domestic demand for consumer goods, leaving the explanatory power of this domestic demand equation unchanged. As we will show further below, this is because the substantial negative income effect is more than offset by a positive substitution effect, leaving the net of the two effects, shown by the regression coefficient, less significant. The regression coefficients on the exchange rate variable clearly suggest an immediate drop in demand for imports of $4.33 billion when the exchange rate drops one point, and an increase in domestic demand of $1.64 billion. However, our confidence in our estimate of this marginal effect remains high, since it is the same as that obtained by subtracting the import regression coefficient from the total consumption coefficient, and both of these are significant. These coefficients will be used below in estimating the total impact on the economy of declines in the exchange rate, including subsequent multiplier effects.
4. The Investment Demand Model

Demand for Investment goods may also decline when exchange rate declines raise import prices, lowering real business income and saving. How much of the decreased demand will be for domestic versus imported investment goods will depend on the marginal propensities to invest (MPI_D or MPI_M) in response to a change in the economy’s real growth rate (i.e., the “accelerator effect”) caused by a declining exchange rate. A secondary decrease in Investment should also occur due to multiplier effects of the original income, reducing savings even further, causing increased crowd out effects.

If investment goods are a “normal” good, the effect on U.S.-produced investment goods should include a positive substitution effect resulting from reduced import demand. The exchange rate coefficient should show the net of this positive effect and also a negative effect on investment of the decrease in real income generated by the exchange rate’s decline. The substitution effect on domestic demand should be positive if investment goods are a “normal” good. The effect should be negative if we find that demand for domestic investment goods declines in favor of imports when real income falls due to a declining exchange rate, even though it means imports are becoming more expensive. Whether the substitution effect is positive or negative and whether the income effect or substitution effect dominate are empirical questions addressed below.

The investment model tested includes key variables traditionally thought to influence investment. See, for example, Jorgenson (1971). Imported investment goods are defined as imported capital goods plus imported industrial supplies and materials. The current period is denoted without a subscript; prior years are subscripted with a -1 or -2. Since the variables in each are the same, the tested equations for domestic (ID) and imported investment (IM) goods all take the general form

\[ \Delta I_D = (\Delta I - \Delta M_{ksm}) = \beta_{D1} \Delta ACC + \beta_{D2} \Delta DEP + \beta_{D3} \Delta CAP_{-1} + \beta_{D4} \Delta T_G - \beta_{D5} \Delta G - \beta_{D6} \Delta r_{-2} + \beta_{D7} \Delta DJ_{-2} + \beta_{D8} \Delta PROF_{-2} + \beta_{D9} \Delta XR_{AV0123} \]  

(7)

[ \Delta I_M = (\Delta M_{ksm}) = \beta_{M1} \Delta ACC + \beta_{M2} \Delta DEP + \beta_{M3} \Delta CAP_{-1} + \beta_{M4} \Delta T_G - \beta_{M5} \Delta G - \beta_{M6} \Delta r_{-2} + \beta_{M7} \Delta DJ_{-2} + \beta_{M8} \Delta PROF_{-2} + \beta_{M9} \Delta XR_{AV0123} \]  

(8)

The variables included in these equations are

- \( \Delta ACC \) = An accelerator variable \( \Delta(Y_t - Y_{t-1}) \)
- \( \Delta DEP \) = Depreciation, a measure of investment needed this year just to replace worn out plant and equipment
- \( \Delta CAP_{-1} \) = A measure of last year’s capacity utilization
- \( \Delta PROF_{-2} \) = A measure of business profitability two years ago

The other variables have the same meanings they had in the consumption equations, with lags as noted.

Our purpose here is not to analyze definitively the components of the investment function, but just to provide estimates of the effect of the exchange rate on investment that have been obtained while controlling for as least some of the other variables that might affect investment, and whose influences might otherwise be picked up by the exchange rate variable due to intercorrelation.

The parameters in this investment demand model were estimated to be:
Total Investment Demand

\[ \Delta I = 0.28 \Delta ACC + 0.95 \Delta DEP + 1.48 \Delta CAP, 1+0.52 \Delta T_G - 0.63 \Delta G - 6.40 \Delta \xi_{-2} - 20 \Delta DJ_{-2} + 16 \Delta PROF, 2 + 6.92 \Delta XR_{AV0123} \]

\[ R^2 = 0.89 \]  

(9)

Demand For Imported Investment Goods

\[ \Delta (M_{km}) = 0.04 \Delta ACC + 0.38 \Delta DEP + 1.52 \Delta CAP, 1 + 0.07 \Delta T_G - 22 \Delta G - 1.54 \Delta \xi_{-2} + 19 \Delta DJ_{-2} - 0.10 \Delta PROF, 2 + 2.52 \Delta XR_{AV0123} \]

\[ R^2 = 0.70 \]  

(10)

Demand For Domestically Produced Investment Goods

\[ \Delta (I-M_{km}) = 0.24 \Delta ACC + 0.57 \Delta DEP - 0.04 \Delta CAP, 1 + 0.44 \Delta T_G - 41 \Delta G - 8.00 \Delta \xi_{-2} - 38 \Delta DJ_{-2} + 26 \Delta PROF, 2 + 4.39 \Delta XR_{AV0123} \]

\[ R^2 = 0.85 \]  

(11)

5. The Exports Demand Model (Using The Nominal Broad XR Index)

There is also an increase in income that occurs because of the increase in exports associated with the decline of the exchange rate. A rough estimate of this effect can be obtained by regressing exports on the 4-year average exchange rate above and the growth in the American GDP over the 1960-2000 period. The income variable serves as a proxy for the growth in our major trading partners’ incomes over this period. Our trading partners’ incomes should have a major effect on the demand for our exports. The results of this regression, using first differences in the data to reduce multicollinearity and stationarity problems, as well as 2SLS, autocorrelation and heteroskedasticity controls are as follows:

\[ \Delta X_0 = 0.08 \Delta (Y_{AV12}) - 5.70 \Delta XR_{AV0123} + 0.42 AR(1) + 0.61 AR(3) \]

\[ R^2 = 53\% \]  

(12)

6. The Tax Growth Model

Part of tax growth is exogenous, i.e., varies with legislative changes in tax rates. However, part is endogenous, i.e., dependent on income growth from year to year. Below we estimate the effect of a change in total income (GDP) on part of tax revenues - the part raised to finance purchases of goods and services. The results of this regression, using first differences in the data to reduce multicollinearity and stationarity problems, as well as 2SLS and heteroskedasticity controls are as follows:

\[ \Delta T_G = 0.26 \Delta Y \]

\[ R^2 = 47\% \]  

(13)
Both the consumption and investment equations above show a positive effect on demand of an increase in tax revenues, presumably by reducing crowd out caused by government deficits. Therefore, in calculating the full effects of a rise in real income due to exchange rate changes, it is important to also measure the secondary boost to income resulting from additional taxes collected as income grows. We might also define tax changes that are government - enacted, i.e., exogenous, as approximately $\Delta T_{EX}$, where

$$\Delta T_{EX} = \Delta T_G - .26 \Delta Y \quad \text{(or)} \quad \Delta T_G = .26 \Delta (Y) + \Delta T_{EX} \quad (14)$$

We say “approximately, because $T_{EX}$ also contains the regression error term.

7. A Model For Calculating Multiplier, Accelerator And Crowd Out Effects Of Exchange Rate Changes

Some readers may be unfamiliar with notation commonly used by economists to denote different parts of the economy, or with commonly used economic terms like “multiplier” or “accelerator”. To illustrate how these terms are used, the following definitions and derivations of the multiplier and accelerator are presented, using simplified versions of our above consumption and investment equations for ease of exposition:

The GDP ($Y$) is comprised of consumer goods ($C$), investment goods ($I$), goods and services produced for the government ($G$) and net exports ($X-M$):

$$Y = C + I + G + (X-M) \quad (15)$$

In a simple model of the economy, demand for consumer goods might be defined as follows

$$C = (c_0 + m_0) + (c_1 + m_{c1})(Y-T_G) + (c_2 + m_{c2}) T_G + (c_3 + m_{c3}) G \quad (16)$$

where ($Y-T_G$) is total income generated producing the GDP minus total taxes; $c_1 + m_{c1}$ are the marginal propensities to consume domestic and imported goods, $T_G$ and $G$ represent the variables measuring the extent to which consumer credit is crowded out by the government deficit. The disaggregated form of the deficit is used ($T_G, G$ separately) instead of just ($T_G - G$) because testing above indicates that the effects of the two variables on crowd out are different.

Demand for investment goods in a simple model of the economy might be described as

$$I = I_0 + (I_1 + m_{i1}) \Delta Y - (I_2 + m_{i2}) r + (I_3 + m_{i3}) T_G + (I_4 + m_{i4}) G \quad (17)$$

where $\Delta Y$ is an “accelerator” variable, indicating $I$ grows (accelerates) in response to the general growth in the economy, $r$ is the real interest rate, $(I_1 + m_{i1})$ are the marginal propensities to purchase domestically produced or imported investment goods in response to a change in the GDP, $(I_2 + m_{i2})$ are the marginal propensities to invest in these goods when interest rates change. $T_G + G$ represent the investment credit crowd out variables, again disaggregated.
Import demand is expressed as

\[ M = M_C + M_I = m_0 + m_{c1} (Y-T) + m_{i1} \Delta Y - m_{i2} r + (m_{c2} + m_{i3}) T_G + (m_{c3} + m_{i4}) G \]  

(18)

i.e., the demand for imported consumer or investment goods is driven by the same variables as is domestic demand.

Substituting (2), (3) and (4) into equation (1) gives

\[ Y = (c_0 + I_0 - m_0) + c_1 (Y - T_G) + I_1 \Delta Y - I_2 r + G + X + (c_2+I_3) T_G + (c_3+I_4) G \]  

(19)

i.e., the domestic GDP is a function of the demand for domestic C,I,G and X goods, as modified by crowed out problems.

Collecting only the Y terms, we get

\[ Y = \left[ \frac{1}{1-c_1} \right] \left[ (c_0 + I_0 - m_0) - c_1 T_G + I_1 \Delta Y - I_2 r + G + X + (c_2+I_3) T_G + (c_3+I_4) G \right] \]  

(20)

where \( \frac{1}{1-c_1} \) is the standard textbook consumption multiplier and equals \( \frac{1}{1-.57} = 2.33 \)

using the marginal propensity to consume domestically produced goods from the domestic consumption regression (Eq. 6) in section 3 above.

However, if we separate \( I_1 \Delta Y \) into its separate components, \( I_1 Y \) and \( - I_1 \Delta Y \), and recollect our current year Y terms, we get a modified multiplier (or multiplier/accelerator) coefficient that combines traditional multiplier and accelerator effects:

\[ Y = \left[ \frac{1}{1-c_1-I_1} \right] \left[ (c_0 + I_0 - m_0) - c_1 T_G + I_1 Y - I_2 r + G + X + (c_2+I_3) T_G + (c_3+I_4) G \right] \]  

(21)

where the numerical value the accelerator/multiplier coefficient is \( \frac{1}{1-.57-.24} = 5.26 \)

using domestic investment regression results above (eq. 11) for the accelerator coefficient. We can further augment this function by noting that the tax component (\( T_G \)) of the “crowd out” variables in both the consumption and investment equation grows as income grows, as shown in our tax growth model above. Also, our consumption and investment regressions above suggest that a rise in taxes depresses consumption spending by decreasing disposable income -$.57B for each billion increase in \( T_G \), but that the same rise in taxes stimulates consumer spending by +$.20B and investment spending by +$.44B by reducing the government deficit crowd out effect, more than offsetting the negative impact of taxes on disposable income, for a net effect of +$.09B. Hence,

\[ (-c_1 + c_2 + I_3) T_G = (-.57 +.20 +.44) T_G \]
\[ = (.07) T_G = (.07) (.26 Y + T_{EX}) \]
\[ = .02 Y +.07 T_{EX} \]  

(using eq. 14 above)
Using this formulation and recombining the Y terms gives a further modified multiplier we will call the “Multiplier/Accelerator/Crowd Out” multiplier (or “M/A/C”):

\[
Y = \frac{1}{(1-c_1-I_1[-c_1+c_2+I_3][.26])} \left[ (c_0 + I_0 - m_0) - c_1 T_{EX} - I_1 Y_{-1} - I_2 r + G + X + (c_2+I_3) T_{EX} +(c_3+I_4) G \right] 
\] (22)

Expressed in first differences this becomes

\[
\Delta Y = \frac{1}{(1-c_1-I_1[-c_1+c_2+I_3][.26])} \left[ - c_1 \Delta T_{EX} - I_1 \Delta Y_{-1} - I_2 \Delta r + \Delta G + \Delta X + (c_2+I_3) \Delta T_{EX} +(c_3+I_4) \Delta G \right] 
\] (23)

where the numerical value of M/A/C multiplier becomes \[
\frac{1}{(1-.57-.24-.02)} = 5.88
\]

This is the multiplier we will use below to calculate the total effect of a change in the exchange rate on U.S. real income.

8. Of Income And Substitution Effects Of A Declining Exchange Rate: Effects On Demand

Does the demand for imports decline when the nominal exchange rate drops one index point? For ‘normal” goods, economic theory suggests both the income and the substitution effects should be negative for imports, each causing a decrease in demand. For normal domestically produced goods, theory suggests the income and substitution effects should work in opposite directions: substitution effects increasing domestic demand, income effects decreasing it, as higher import prices cause people to shift from imports to domestic goods.

Our statistical results for consumption demand are consistent with this theory. Heim (2009) suggests a method for separating the income and substitution effects of changes in nominal exchange rates on consumption. The method involved shows that the “pure” income effect must be the same for both domestically produced and imported consumer goods. It also shows the same is true for substitution effects, except for sign. Once this is done, it is relatively simple, using regression coefficients as estimates of the total effect of exchange rate changes, to calculate the separate income and substitution effects. Application of this method to consumption yields the results given in Table 1 below, where regression coefficients on the exchange rate variable equal total effects, and clearly indicating consumer imports are normal goods, and where substitution effects clearly dominate.

<table>
<thead>
<tr>
<th></th>
<th>Domestic Goods</th>
<th>Imports</th>
</tr>
</thead>
<tbody>
<tr>
<td>Income Effect</td>
<td>- $ 1.345 Billion</td>
<td>- $ 1.345 Billion</td>
</tr>
<tr>
<td>Substitution Effect</td>
<td>+ $ 2.985 Billion</td>
<td>- $ 2.985 Billion</td>
</tr>
<tr>
<td>Total Effect</td>
<td>+ $ 1.640 Billion</td>
<td>- $ 4.330 Billion</td>
</tr>
</tbody>
</table>

Table 1: Consumer Goods Income and Substitution Effects
However, application of this method to investment indicates investment imports may be “inferior” goods, i.e., goods the U.S. substitutes out of as real income increases, and into as real income decreases. Applying the method yields the results indicated in Table 2 below, where regression coefficients on the exchange rate variable equal total effects, the signs on the substitution effects clearly suggests investment imports are inferior goods, and where income effects clearly dominate.

These results show the initial effects on demand for consumer and investment goods of a decline in the exchange rate, ceteris paribus. However, since every change in consumption or investment changes the GDP, which in turn again changes consumption and investment, etc., there will also be multiplier effects. The initial effect on domestically produced consumer goods (+$1.64B), plus the initial effect on domestically produced investment goods (-$4.39B), plus the initial effect on U.S. exports (+ $5.70B) will be subject to multiplier effects. This multiplier used (5.88) will be the M/A/C multiplier developed in Section 7 (eq. 23) above.

<table>
<thead>
<tr>
<th>Table 2: Investment Goods Income and Substitution Effects</th>
</tr>
</thead>
<tbody>
<tr>
<td>Domestic Goods</td>
</tr>
<tr>
<td>Income Effect</td>
</tr>
<tr>
<td>Substitution Effect</td>
</tr>
<tr>
<td>Total Effect</td>
</tr>
</tbody>
</table>

The results above hang on the accuracy of our regression coefficient estimates of the effects on consumption and investment of a change in the exchange rate.

The regression coefficient on the exchange rate variable in the consumer imports equations is statistically significant, and this coefficient is important in our analysis below. The coefficients on the variables in the domestic consumption and total consumption equations are not statistically significant. However, they are the same as that obtained by subtraction of the statistically significant estimate for imports from either the total consumption or domestic consumption coefficients. Hence, these coefficients may be reasonable for use in estimating how exchange rate changes affect the demand for consumer goods. For the three investment equations, two of the three exchange rate coefficients are significant at the 5% level or better: total investment and imports investment. In addition, the exchange rate coefficient for domestic investment goods is statistically significant at the 10% level, and its value is exactly the same as is obtained by subtraction of the two significant estimates. Hence, we feel reasonably confident in the internal consistency of all three of our point estimates.

We do not know with any certainty how much prices change when the exchange rate changes. A one point decrease the exchange rate (which would represent an eight tenths of one percent decrease from the rate’s 119.45 level in the year 2000) could increase import prices by 0.8% percent, if the exchange rate change was passed entirely through to the consumer. However, recent evaluation by Federal Reserve staff of the “pass through” of exchange rate changes to import prices 1985 -2005 suggests that import prices only change about half as much as the exchange rate changes, and not less than that as some others have recently claimed (Hellerstein, Daly & Marsh, 2006). This study finds a change in combined consumer and investment demand, $9.6 billion, equal to about ¾ of dollar value of full 100% pass through. In
the year 2000, U.S. total real imports (1996 dollars) were $1,532 billion. A 0.8% decrease in the 
exchange rate, if fully passed through, would be expected to increase import costs by 0.8%, or 
$12.26 billion, decreasing real incomes in the U.S. by the same amount. Real disposable income 
also decreases the same amount, since there is no tax effect: nominal income (on which taxes are 
based) remains the same. But again, these conclusions would be based on our full pass through 
assumption.

Our approach is different. We make no explicit attempt to calculate the exact change in 
import prices resulting from changes in the exchange rate. Rather, we carefully measure how 
much consumption and investment vary when the exchange rate changes, controlling for other 
factors likely to cause changes in these two variables. We say, in essence, However much a one 
point decline in exchange rates changes in import prices, the change causes a +1.64 billion 
change in demand for domestic consumer goods, a $4.39 billion decline in demand domestic 
investment goods, and a $5.70 change in demand for exports.

9. Three Methods For Calculating The Impact On The Gdp Of A Change In The 
Exchange Rate

Three separate methods, all yielding the same results, are used to compute the effect of a 
change in the exchange rate on the GDP (Y):

Method 1: Use marginal effects estimates from the above domestic 
investment, consumption and export regressions to estimate the initial change in 
domestic consumption, investment and the GDP resulting from a one index point 
drop in the trade weighted exchange rate. Apply the M/A/C multiplier (5.88) to the 
result

Method 2: Use the method favored in many large scale econometric models 
of the economy (Fair 2003, Pindyck & Rubinfeld 1991). This involves separately 
estimating $\Delta C_D$, $\Delta I_D$, $\Delta G$ and $\Delta X$ (using the equations above), and simply summing 
the results to get $\Delta Y$

Method 3: Formally Construct a Keynesian IS curve, and predict $\Delta Y$ from its 
determinants and the multiplier implied by the function. It is a slightly more formal 
presentation of Method 1.

Each of the methods can serve as a check on the estimates obtained from the others.

9.1 Method 1

$ + 1.64B$ (Billion) - Total Estimated Effect (Positive Substitution Minus Negative Income 
Effect) Of A One Point Exchange Rate Decline On Demand For 
Domestically Produced Consumer Goods ($C_D$)

$ - 4.39B$ - Total Effect (Negative Substitution plus Negative Income Effect) Of 
Demand Decline For Domestically Produced Investment Goods ($I_D$)

$ + 5.70B$ - Increase In Exports ($X$)

$ + 2.95B$ - Initial Net Decline in Real U.S. Income from Impact of Declining 
Exchange Rate:

$ \times 5.88$ - Multiplier/Accel/Crowd Out (M/A/C)Effect

$ + 17.35B$ - Decline in Real Income ($Y$) after (MAC) Multiplier Effects
- 4.51B  - Δ Income - Related Taxes Calculated At Historic .26 Rate (.26*17.35 = 4.51B)

$ +12.84B  - Δ(Y-T_G) = Decline In Disposable Income Associated With A One Point Decline In The Exchange Rate

To see the impact of decreased credit availability (crowd out) due to decreased tax collections:

+ 0.90B = ΔC_D Due to Reduced Crowd Out From Increased Taxes = (.20)($ +4.51B)
+ 1.22B = ΔC_M Due to Reduced Crowd Out From Increased Taxes = (.27)($ +4.51B)

With this information we can summarize the changes in consumption and saving resulting from the increase in disposable income of $ 12.48B as follows:

$+12.84B  Δ(Y-T_G) $+12.84B  Δ(Y-T_G) $+12.84B  Δ(Y-T_G)
$ +0.90B  ΔC_D $+ 0.77B  ΔC_M $+ 4.75B  ΔSavings
$ +1.22B  ΔC_M $  $ - 2.34B  ΔC_M

9.2 Method 2:

We repeat the above investment and consumption regression equations for easy reference as we calculate Method 2:

ΔC_D = Δ(C- M_{km,sm}) = .57Δ(Y-T_G) + .20ΔT_G0 + .24 ΔG0 - 2.28 ΔPR0 - .34ΔDJ - 1.64 ΔXR_{AV0123}
ΔC_M = Δ(M_{km,sm}) = .06Δ(Y-T_G) + .27ΔT_G0 - .18 ΔG0 - 3.94 ΔPR0 + .26 ΔDJ + 4.33 ΔXR_{AV0123}

ΔI_D = Δ(I-M_{km}) = .24ΔACC +.57ΔDEP -.04ΔCAP +.44 ΔT_G - .41ΔG - 8.00Δr - .38 ΔDJ + .26 ΔPROF + 4.39ΔXR_{AV0123}
ΔI_M = Δ(M_{km}) = .04ΔACC +.38ΔDEP +1.52ΔCAP +.07 ΔT_G - 223ΔG + 1.54Δr +19ΔDJ - .10 ΔPROF + 2.52 ΔXR_{AV0123}

From these equations, we see three variables through which investment is affected by changes in the exchange rate:

1. the decrease in the accelerator income variable in the investment equation, or the disposable income variable in the consumption equation, due to the decrease in gross real income caused by the one point decline in XR_{AV0123}
2. the decline in tax collections because of the decline in real income caused by the increase in import prices, and
3. through the one point change in the exchange rate variable

In this case then, the estimated decline in domestic investment will be

$$\Delta I_D = \Delta(I-M_{km}) = .24ΔACC +.44 ΔT_G + 4.39 ΔXR_{AV0123}$$
$$= (.24)(+$17.35B) + (.44)(+$4.51) + (4.39B)(-1)$$
$$= +4.16 + 1.98 - 4.39$$

= $ +1.75B
where the change in taxes $\Delta T_G$ is the difference between the change in gross income ($\Delta Y$) and the change in disposable income $\Delta(Y-\Delta T_G)$ given above.

We can also estimate the decrease in demand for imported investment goods as

$$\Delta I_M = \Delta(M_{ksm}) = 0.04 \Delta ACC + 0.07 \Delta T_G + 2.52 \Delta XR_{AV0123}$$

$$= (0.04)(\$ +17.35B) + (0.07)(\$ +4.51B) + (2.52)(-1)$$

$$= \$ - 0.69B + \$ 0.32B - 2.52$$

$$= \$ - 1.51B$$

By similar reasoning, we see that the changes in the demand for domestic and imported consumer goods are as follows

$$\Delta C_D = \Delta(C-M_{m-ksm}) = 0.57 \Delta(Y-T_G) + 0.20 \Delta T_G - (1.64) \Delta XR_{AV0123}$$

$$= (0.57)(\$ +12.84) + (0.20)(\$ +4.51 ) - (1.64)(-1)$$

$$= + 7.32B + 0.90B + 1.64$$

$$= \$ + 9.86B$$ (same result as method 1)

and

$$\Delta C_M = \Delta(M_{m-ksm}) = 0.06 \Delta(Y-T_G) + 0.27 \Delta T_G + 4.33 \Delta XR_{AV0123}$$

$$= (0.06)(\$ +12.84) + (0.27)(\$ +4.51B) + 4.33(-1)$$

$$= + 0.77B + 1.22B - 4.33$$

$$= \$ -2.34B$$ (same result as method 1)

So, by Method 2 we have

$$\Delta Y = \Delta C_D + \Delta I_D + \Delta G + \Delta X$$

$$= \$ +9.86 +1.75 + 0 + 5.70$$

$$= \$ + 17.31B$$ (Same result as Method 1, except for rounding)

9.3 Method 3:

Using the formal Keynesian “IS” curve method for calculating the GDP shown in Section 7 above:

$$\Delta Y = \Delta C_D + \Delta I_D + \Delta G + \Delta X$$

$$= (0.57\Delta(Y-T_G) + 0.27\Delta T_G - 1.64\Delta XR_{AV0123}) + (0.24\Delta ACC + 0.44\Delta T_G + 4.39\Delta XR_{AV0123}) + 0 + 5.70 \Delta XR_{AV0123}$$

$$= (5.88)(+1.64 - 4.39 + 5.70) + (5.88)(0.07)(\Delta T_{EX})$$

$$= (5.88)(+2.95B) + 0$$

$$= \$ + 17.35 B$$ (Same result as by Methods 1 and 2)

10. Exchange Rate Effects On The Trade Deficit And U.S. Assets

The estimated decline in the U.S. trade deficit resulting from a one point decline in the exchange rate is the sum of the resulting decrease in imports and the increase in exports

$ 2.34B - \text{Decline in } C_M$

$ 1.51B - \text{Decline in } I_M$
The annual increase in savings of $4.75B is the estimate the growth in domestic assets due to the exchange rate decline, which causes an increase in real income and saving. This growth in assets can be coupled with the annual decrease in U.S. assets required to fund the trade deficit ($9.55B). The two together mean U.S ownership of (U.S.) assets increases an estimated $14.30 billion annually when the exchange rate declines one point.

Our reasoning is as follows: every trade deficit is financed by a transfer of ownership of domestically owned assets (including money), or claims to them, to other countries or their citizens. This is how the money is raised that allows one country to buy more goods from another than the other wants to buy from the first. A decline in the trade deficit reduces the amount of domestically – owned assets that have to be sold or borrowed against each year to finance the deficit. There is also an increase in domestic savings (i.e., a growth in U.S. – owned assets) that occurs because the exchange rate drops. This increase in savings results from income increase caused by rising demand for domestic consumer, investment and export goods as a result of the exchange rate decline. Putting both factors together, the study’s estimate that a one point reduction in the exchange rate would prevent a ($9.55B 4.75B = $14.30B) decrease in U.S. ownership of U.S. capital assets each year that otherwise would occur.

Conversely, a rise in the U.S. exchange rate of one point would give the same results as above, but with the opposite sign. Income would decrease $17.35B, disposable income by $12.84B, the trade deficit would rise by $9.55B and savings would fall $4.75B, etc. This would reduce the annual net growth of U.S. assets by $14.30 billion.

11. Conclusions

The analysis above indicates that when the nominal trade-weighted exchange rate index falls by one point, the results are as follows:

1. From Method #2 above we have
   a. a decrease in demand for imported consumer and investment goods and services estimated at $3.85B (2.34B C_M, + 1.51B I_M).
   b. an increase in demand for domestically produced consumer and investment goods and services of $17.31B = ($9.86B C_D, + 1.75B I_D + $5.70B X). Our study is too macroeconomic in nature to be able to say which specific industries are affected
2. The trade deficit would likely decrease an estimated $9.55 billion, due to the $3.85B reduction in imported consumer and investment goods, and $5.70B increase in exports.
3. Because a one point drop in the exchange rate is small, the trade deficit as a percent of GDP would only decline one tenth of a percent from 4.3% to 4.2%, if we measure the decline using 2000 as the base year:
4. This suggests that even a fairly large, say 10%, drop in the exchange rate would only decrease the trade deficit as a percent of GDP moderately, by 1.1 percent, from 4.3 to 3.2%, or $95.5B, using the year 2000 GDP and trade deficit as the base year against which to measure the change, as shown in Table 3 below. This decrease in the trade deficit would be accompanied by a small 1.9% increase in the GDP ($173.5B) in year 2000 dollars. Using the numbers from Method 2, multiplied by 10, we have:

$$\Delta Y = \Delta C_D + \Delta I_D + \Delta G + \Delta X$$

$$-173.1B \approx +98.60B +17.50B + 0 +57.0B$$

<table>
<thead>
<tr>
<th>Table 3: Exchange Rate Impact on GDP and Trade Balance</th>
</tr>
</thead>
<tbody>
<tr>
<td>Real GDP</td>
</tr>
<tr>
<td>Actual 2000 Data</td>
</tr>
<tr>
<td>Effect of 1Pt Drop XR</td>
</tr>
<tr>
<td>Effect of 10Pt Drop XR</td>
</tr>
<tr>
<td>Effect of 19.6Pt Dr XR</td>
</tr>
</tbody>
</table>

However, in the period 2000 – 2008, The U.S. exchange rate dropped even more significantly. The Nominal Broad Index dropped 19.6 points (16.4%), from 119.45 to 99.83

Using the 19.6 point drop in the Nominal Broad Index during the 2000-08 period, suggests that this would have been associated with a decrease in the trade deficit as a percent of GDP by 2.1 percentage points, from 4.3% to 2.2%, or $187.2B, had other factors affecting the deficit, like U.S. income and wealth growth for all other reasons remained constant at 2000 levels. This decrease in the trade deficit would have been accompanied by a 3.7% increase in the GDP or $340.0B in year 2000 dollars over the eight years (other things equal). Using the numbers from Method 2, multiplied by 19.6, we have:

$$\Delta Y = \Delta C_D + \Delta I_D + \Delta G + \Delta X$$

$$+340.0 \approx +193.26 +34.30B + 0 +111.72B$$

The $340B increase in GDP associated with the estimated 19.6 point (or 16.4%) 2000-08 decline in exchange rates, would have resulted in a 3.7% increase in 2000 - level real GDP, ceteris paribus. However, Bureau of Economic Analysis data indicated the real GDP grew 18.7% during this period. Presumably, had the exchange rate decline not occurred, it would have grown 3.7% less, decreasing the average growth rate during the period from 2.3% to 1.9% - nearly 1/2 of a percent per year. This suggests the actual annual growth rate was actually higher than it might have been, had the decline in the exchange rate not occurred. Thus, the evidence indicates that the cheaper dollar of the 2000-2008 period did stimulate the U.S. economy.

5. The $9.55 decline in the U.S. Trade deficit associated with a one point drop in the nominal Broad exchange rate index reduces the need for annual transfers of U.S. financial or real assets (including dollars) to foreign ownership. Of course, other transfers are still needed to
pay for the remaining annual trade deficit. This means a decline in the amount of U.S.
owned assets that have to be transferred to the rest of the world to pay for our current demand
for more goods and services from other countries than they want to buy from us. Coupled
with the increase in U.S. savings associated with the decline of the exchange rate, we
estimate each point decline increases U.S. ownership of assets $14.30 billion.

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